# Bayesian Analysis

# the Generalized Linear Model

in

站长最材 sc.chinaz.com **Group member:** 



陈文琳

# **Our goals**



provide enough guidance so we can combine generalized linear models with the ideas of Bayesian analysis.





#### **Classical linear regression model**

$$Y = X\beta + \varepsilon$$

Where we obtain the following variables

- Y endogenous variable
- X exogenous variable
- $\beta$  the regression coefficient
- $\mathcal{E}$  —— the random error

#### A possible assumption of $\ {\cal E}$

$$\varepsilon \sim N(0, \sigma^2)$$

## We denote that $\mu = X\beta$

$$\Rightarrow \mu = E(Y | X)$$

**Several properties of linear regression model** 

$$\star E(Y) = \mu = X\beta$$

$$\star X, Y$$
 are continuous variables;

$$\star \quad Y \sim N(\mu, \sigma^2)$$

#### shortage of linear regression model

The assumption that Y have normal distribution is impractical

Exogenous variable can only effect on endogenous variable through addition

Endogenous variable must be continuous variable

discontinuous characters



Counted variables

0-1 variables  $z \sim P(\lambda)$ 



variables do no follow the normal distribution

variables follow Gamma distribution , binomial distribution

And so on

#### **Difference between linear model and GLM**

1.Endogenous variables can follow any distribution in exponential family of distributions ;

2. We induct link function  $\theta(\mu) = X\beta$  to measure the mean of endogenous

#### **Several properties of linear regression model**

x, y can be continuous or discrete variables

$$\star f(y,\mu,\phi) = \exp\left[\frac{y(\mu) - b(\mu)}{a(\phi)} + c(y,\phi)\right]$$

$$\star \quad E(y) = \mu, \theta(\mu) = X\beta$$

A generalized linear model is specified in three stages:

1. The linear predictor  $\theta = X\beta$ 

2. The link function  $g(\cdot)$  that relates the linear predictor to the mean of the outcome

variable: 
$$\mu = \theta^{-1}(X\beta)$$

3. The random component specifying the distribution of the outcome variable y with mean

 $E(y|X) = \mu$ . The distribution can also depend on a dispersion parameter,  $\varphi$ .

The exponential family

The probability density function of the exponential family

$$f(y|\theta, \Phi) = \exp\{\frac{y\theta - b(\theta)}{\Phi} + c(y, \Phi)\}$$

 $\theta$  is a canonical parameter and a function of the mean of the outcome variable  $~\mu$ 

 $b(\theta)$  is a function of  $\theta$  and has positive second order derivative

 $\Phi$  is a dispersion parameter that plays a role in defining the variance of y  $c(y, \Phi)$  is a function of y and  $\Phi$ 

Link function , expectation and variance of y?

We can use normal distribution  $f(y|\mu, \sigma^2)$  to find connections with the exponential family.

$$\mu = \theta$$
,  $b(\theta) = \frac{\theta^2}{2}$ ,  $\Phi = \sigma^2$ 

More generally, we can set  $\theta$  as linear predictor and get link function:

$$\theta(\mu) = X\beta$$

Analogously, we can get the expectation and variance of y:

$$E(y) = b'(\theta), \quad Var(y) = b''(\theta)\Phi$$

### Likelihoods

The mean of the distribution of y, given X, is determined by X $\beta$ : E(y | X) =  $\theta^{-1}(X\beta)$ .We use the same notation as in linear regression whenever possible, so that X is the n  $\times$  p matrix of explanatory variables and  $\theta = X\beta$  is the vector of n linear predictor values.

If we denote the linear predictor for the ith case by  $X_i\beta$  and the variance or dispersion parameter (if present) by  $\Phi$ , then the data distribution takes the form

$$p(y|X,\beta,\Phi) = \prod_{i=1}^{n} p(y_i|X_i\beta,\Phi)$$

## Standard GLM likelihoods

Normal distribution

Normal distribution has an identity link function  $\theta(\mu) = \mu$ .

With a assumption  $\varepsilon \sim N(0, \sigma^2)$ , we can learn that linear regression model  $y = X\beta + \varepsilon$  is a special case of the generalized linear model, for

 $y|\beta, X, \sigma^2 \sim N(X\beta, \sigma^2).$ 

Standard GLM likelihoods

#### Poisson distribution

Counted data are often modeled using a Poisson model. The Poisson generalized linear model, often called the Poisson regression model, assumes that y is Poisson with mean  $\mu$  (and therefore variance  $\mu$ ).

The link function is typically chosen to be the logarithm, so that  $\log \mu = X\beta$ . The distribution for data  $y = (y_1, \dots, y_n)$  is thus

$$p(\mathbf{y}|\boldsymbol{\beta}) = \prod_{i=1}^{n} \frac{1}{y_i!} e^{-\exp(\theta_i)} (\exp(\theta_i))^{y_i}$$

where  $\theta_i = Xi\beta$  is the linear predictor for the i-th case.

## Standard GLM likelihoods

#### Binomial distribution

Suppose that  $y_i \sim \text{Bin}(n_i, \mu_i)$  with  $n_i$  known. It is common to specify the model in terms of the mean of the proportions  $y_i / n_i$ , rather than the mean of  $y_i$ . Choosing the logit transformation of the probability of success,  $\theta(\mu_i) = \log(\mu_i/(1-\mu_i))$ , as the link function leads to the logistic regression model.

The distribution for data y is

$$\mathbf{p}(\mathbf{y}|\boldsymbol{\beta}) = \prod_{i=1}^{n} \binom{n_i}{y_i} \left(\frac{e^{\theta_i}}{1+e^{\theta_i}}\right)^{y_i} \left(\frac{1}{1+e^{\theta_i}}\right)^{n_i - y_i}$$

The combination of GLM and Bayesian Analysis

GLM: nonlinear model; Discrete data

Bayesian Analysis: Small sample size Too many parameters

. . .

#### reserve for outstanding losses

To fail to draw or carry down various kinds of liability reserves or fail to draw reserve for outstanding losses according to the provision of this law. In general, we take the unbiased estimator of expected value of outstanding losses as the reserve for outstanding losses.

## Triangular flow

事故年 (i)	进展年())					
	1	2	3	4		п
1	Gu	G12	Gis	G14		С. л
2	C21	G22	C23	G24		
3	Gi	G12	C53			
n	Cns					

#### We define the following variables

- $X_i$  the aggregate amount of outstanding losses in the ith year
- $\mathcal{Y}_{j}$  the probability of outstanding losses in the jth year
- $C_{ij}$  the expected amount of outstanding losses in the jth year when one insured in the ith year

**Over-dispersed Poisson model** 

 $z \sim p(\lambda)$  $x = \varphi z, (\varphi \ge 1)$ 

The prior distribution

$$x_j \sim Gamma(\alpha_i, \beta_i)$$

#### Then we obtain that

$$f(x, y | c_{ij}, i = 1, \dots, n, j = 1, \dots, n - i + 1, \varphi) \propto \prod_{i=1}^{n} \prod_{j=1}^{n-i+1} f(c_{ij} | x, y, \varphi) \prod_{i=1}^{n} f(x_i) f(y_i)$$

Therefore, we can measure the expected amount of outstanding losses in the jth year when one insured in the ith year

$$\hat{c}_{ij} = E(c_{ij}) = \exp(\hat{\eta}_{ij})$$

We can measure the reserve for outstanding losses

$$R = \sum c_{ij}$$

